doi: 10.1093/bja/aex138 Review Article

REVIEW ARTICLE

Perioperative goal-directed therapy with uncalibrated pulse contour methods: impact on fluid management and postoperative outcome

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Abstract

Previous meta-analyses suggest that perioperative goal-directed therapy (GDT) is useful to decrease postoperative morbidity. Most GDT studies analysed were done with pulmonary artery catheters, oesophageal Doppler and calibrated pulse contour methods. Uncalibrated pulse contour (uPC) techniques are an appealing alternative but their accuracy has been questioned. The effects of GDT on fluid management (volumes and volume variability) remain unclear. We performed a meta-analysis of randomized controlled trials investigating the effects of GDT with uPC methods on postoperative outcome. The primary endpoint was postoperative morbidity. Fluid volumes and fluid volume variability (standard deviation/mean) over the GDT period were also studied. Nineteen studies met the inclusion criteria (2159 patients). Postoperative morbidity was reduced with GDT (OR 0.46, 95% CI 0.30–0.70, P<0.001). The volume of colloids was higher [weighted mean difference (WMD) +345 ml, 95% CI 148–541 ml, P<0.001] and the volume of crystalloids was lower (WMD –429 ml, 95% CI –634 to –224 ml, P<0.01) in the GDT group than in the control group. However, the total volume of fluid (WMD –220 ml, 95% CI –590 to 150 ml, P=0.25) and the variability of fluid volume (34% vs 33%, P=0.98) were not affected by GDT. The use of GDT with uPC techniques was associated with a decrease in postoperative morbidity. It was not associated with an increase in total fluid volume nor with a decrease in fluid volume variability.

Key words: cardiac output; general surgery; haemodynamics

Many studies suggest that perioperative goal-directed therapy (GDT) is useful to decrease postoperative morbidity, hospital length of stay and hospital costs.^{1–3} As a result, in patients undergoing major surgery, the use of GDT is now recommended by several guidelines and consensus statements from international experts.^{4–7} The first perioperative GDT studies were done 20–30 yr ago with the pulmonary artery catheter.^{8–10} Then, other studies followed where haemodynamic parameters were

derived from the oesophageal Doppler¹¹¹² or from calibrated pulse contour methods.¹³¹⁴ Uncalibrated pulse contour (uPC) methods are relatively new in the GDT arsenal since they became available only a decade ago.¹⁵ They are quick to set up, easy to use, not operator dependent, not affected by electrocautery and are increasingly used for haemodynamic monitoring during major surgery.¹⁶ However, their accuracy and precision have been questioned when compared with clinical

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reference methods, such as thermodilution and echocardiography.^{17–20} Whether uPC techniques can be useful to guide haemodynamic therapy and improve post-surgical outcome has been investigated by several randomized controlled trials (RCTs) yielding conflicting results.

Both insufficient and excessive fluid administration are associated with an increase in postoperative complications.^{21 22} Optimizing haemodynamic parameters such as stroke volume and cardiac output with fluid may result, at least in theory, in excessive fluid administration. In addition, recent studies have reported a very large variability in the volume of fluid administered to surgical patients during the perioperative period.^{22 23} By analogy with manufacturing and the Six Sigma concept, it has been suggested that variability of clinical practices is the enemy of quality of care,²⁴ and that the beneficial effects of GDT may be related to the harmonization of fluid management.^{25 26} Therefore, we performed a meta-analysis of RCTs to clarify the impact of GDT with uPC methods on postoperative morbidity, on fluid volume and on fluid volume variability.

Methods

Eligibility criteria

According to Preferred Reporting Items for Systematic reviews and Meta-Analyses (PRISMA), studies were searched using the following eligibility criteria.²⁷ Participants were adult (age 18 yr or over) patients undergoing elective or emergency surgery. Studies involving mixed population of critically ill or nonsurgical patients were excluded. The intervention was defined as GDT with uPC methods. RCTs comparing the effects of GDT vs standard or usual fluid management were considered for analysis. No language (i.e. article in English), publication date or publication status restrictions were imposed when selecting the studies to be analysed. Primary outcome measure was postsurgical morbidity, defined as the proportion of patients developing one or more post-surgical complications. Post-surgical infectious, cardiac, respiratory, renal and abdominal complications, as well as hospital length of stay and mortality, were assessed as secondary outcome variables. Abdominal complications included both gastro-intestinal and liver complications. The volume of crystalloids and of colloids, as well as the total volume of fluid received during the GDT period were also analysed.

Information sources

Various search strategies were performed to retrieve relevant studies by using MEDLINE, the Cochrane Library and EMBASE databases (last update January, 2016). No date restriction was applied for MEDLINE and The Cochrane Library databases whereas the search was limited to 2006–16 for the EMBASE database. Additional trials were searched in the DARE database and the reference lists of previously published reviews and retrieved articles.

Search

We used the following terms to search for studies: randomized controlled trial, controlled clinical trial, goal directed, goal oriented, goal target, cardiac output, cardiac index, oxygen delivery, oxygen consumption, cardiac volume, stroke volume, fluid therapy, fluid loading, fluid administration, optimization, optimisation, pulse pressure variation, pleth variability index, stroke volume variation, systolic pressure variation (see Supplementary data S1 for details regarding the search strategy).

Study selection

Two investigators (N.B., M.T.G.) first examined each title and abstract to identify potentially relevant articles. The eligibility of the retrieved full-text articles was independently determined by two investigators (N.B., F.M.). The analysis was limited to trials done with uPC methods.

Data collection process

Data were independently collected by two investigators (M.T.G., F.M.) with any discrepancy resolved by re-inspection of the original article. To avoid transcription errors, the data were inputted into statistical software and re-checked by a third investigator (N.B.).

Data items

Data abstraction included type of surgery, number of patients, type of uPC method, GDT protocol end-points, postoperative morbidity, complications, mortality and hospital length of stay. The volume of colloid and crystalloid solutions administered during the GDT period was also collected. When information was not found in original manuscripts, authors were contacted to maximize the number of data available for analysis.

Risk of bias in individual studies

A domain-based evaluation, as proposed by the Cochrane Collaboration,²⁸ was used to evaluate the methodological quality of RCTs. This is a two-part tool, addressing seven specific domains (sequence generation, allocation concealment, blinding of participants and personnel, blinding of outcome assessment, incomplete outcome data, selective outcome reporting and 'other issues') that are strongly associated with bias reduction.^{29 30} Each domain in the tool includes one or more specific entries in a 'Risk of bias' table. Within each entry, the first part of the tool describes what is reported in the study, in sufficient detail to support a judgment about the risk of bias. The second part of the tool assigns a judgment relating to the risk of bias for that entry. This is achieved by assigning a judgment of 'Low risk', 'High risk' or 'Unclear risk' of bias. After each domain was completed, a 'Risk of bias summary' table was generated. The green symbol plus indicates low risk of bias, the red minus symbol indicates high risk of bias and the white colour indicates unclear risk of bias. For each study, the total number of green plus symbols was calculated: trials with five or six green plus symbols were considered as having an overall low risk of bias. With regard to blinding, studies in which the outcome variables were collected by investigators not aware of the GDT strategy were considered adequately masked.

Summary measures and planned method of analysis

Meta-analytic techniques (analysis software RevMan, version 5.3 Cochrane Collaboration, Oxford, England, UK) were used to combine studies using Odds Ratios (OR) and 95% confidence intervals (CI) for dichotomous variables, and weighted mean difference (WMD) and 95% CI for continuous variables. A statistical difference between groups was considered to occur if the pooled 95% CI did not include 1 for the OR. An OR <1 favoured GDT when compared with standard haemodynamic treatment. Two-sided P-values were calculated. Statistical heterogeneity and inconsistency were assessed by using the Q and I^2 tests, respectively. When the P-value of the Q test was <0.10 and/or the I² was >40%, heterogeneity and inconsistency were considered significant. Both random-effects and fixed-effects models were used for analyses. Random-effects model results are presented in the abstract and figures whereas fixed-effects model are reported and compared in the supplementary data. For the primary outcome (postoperative morbidity), a sensitivity analysis was performed by focusing on studies with a low risk of bias, on studies where cardiac output was used as a target parameter, and on studies where the total fluid volume was <5 litres. A trial sequential analysis was also performed to adjust for random error risk. We calculated information size and monitoring boundary anticipating a 30% relative risk reduction in postoperative morbidity with GDT. We set risk of type I at 5% and power at 80%.

Coefficients of variability of fluid volume were calculated as standard deviation/mean, expressed as a percentage, and compared using a Student's t-test. When mean and standard deviation values were not reported in original manuscripts and not provided when authors were contacted, they were estimated from median and inter-quartile values as previously described.³¹

Results

Study selection

The search strategy identified 3312 (MEDLINE), 9858 (Cochrane Library) and 2212 (EMBASE) articles. Fourteen articles were identified from the reference list of other articles. After initial screening and subsequent selection, a pool of 102 potentially relevant trials was identified. The subsequent eligibility process excluded 83 articles. Overall, 19 articles with a total sample of 2159 patients, were considered for analysis (see Supplementary data Figure S2).

Study characteristics

All studies were published between 2008 and 2015^{32–49} (Table 1). The risk of bias assessment revealed that eight out of 19 studies were considered as having a low risk of bias (see Supplementary data Table S3).

Postoperative outcome

The number of patients with one or more postoperative complications was significantly reduced by GDT, when using both a random-effects (Fig. 1) or a fixed-effects model analysis (see Supplementary data Table S4). It also decreased significantly in the GDT group when focusing on studies with a low risk of bias, on studies where cardiac output was used as a target variable, and on studies where the total volume of fluid was <5 litres (see Supplementary data Table S5). In the trial sequential analysis, the cumulative z curve crossed the monitoring boundary before reaching the information size, indicating firm evidence that GDT reduced postoperative morbidity⁵⁰ ⁵¹ (Fig. 2).

A significant reduction in infectious, cardiac and abdominal complications was observed in favour of GDT, when using both the random-effects and the fixed-effects models (see Supplementary data Table S4). Renal and respiratory complications, as well as mortality, were not significantly reduced by GDT (see Supplementary data Table S4). Hospital length of stay tended to decrease (-1.35 days, CI -2.78 to -0.08 days, P=0.06).

Fluid management

Patients in the GDT group received more colloid (Fig. 3) and less crystalloid (Fig. 4) than patients from the control group. The total volume of fluid was not significantly different between the GDT and the control group when using the random-effects model (Fig. 5), and was significantly lower in the GDT group

Table 1 Characteristics of studies analysed. HIPEC, hyperthermic intraperitoneal chemotherapy; AAA, abdominal aortic aneurysm; LiDCOr, LiDCO rapid; SVV, stroke volume variation in %; PPV, pulse pressure variation in %; SVI, stroke volume index in ml m^{-2} ; CI, cardiac index in l min m^{-2} ; SVplateau, stroke volume corresponding to the plateau of the Frank–Starling relationship; DO₂I, oxygen delivery index in ml min⁻¹ m⁻²; ScvO₂, central venous oxygen saturation in % SVR, systemic vascular resistance; CO, cardiac output

First author (ref)	Surgery, n	uPC method	Haemodynamic goals	
Benes ³²	Abdominal and vascular, 120	FloTrac	SVV<10, CI>2.5	
Cecconi ³³	Hip, 40	FloTrac	SVplateau, $DO_2I > 600$	
Colantonio ³⁴	HIPEC, 80	FloTrac	SVV<15, SVI>35, CI>2.5	
Correa-Gallego ³⁵	Liver, 135	FloTrac	SVV<15	
Funk ³⁶	AAA, 40	FloTrac	SVV<13, CI > 2.2	
Hand ³⁷	Head and neck, 94	FloTrac	SVV<13, CI>3, SVR>800	
Kapoor ³⁸	Cardiac, 27	FloTrac	SVV<10, SVI>30, CI>2.5, ScvO ₂ >70, DO ₂ I>450	
Lai ³⁹	Abdominal, 220	LiDCOr	SVV<10	
Mayer ⁴⁰	Abdominal, 60	FloTrac	SVV<12, SVI>35, CI>2.5	
Pearse ²	Abdominal, 734	LiDCOr	SVplateau	
Peng ⁴¹	Orthopaedic, 80	FloTrac	SVV<10 (supine) or < 14 (prone)	
Poso ⁴²	Bariatric, 50	FloTrac	SVV<12, SV and CO $>$ 70% of baseline	
Ramsingh ⁴³	Abdominal, 38	FloTrac	SVV<12	
Salzwedel ⁴⁴	Abdominal, 160	ProAQT	PPV<10, CI > 2.5	
Scheeren ⁴⁵	Abdominal, 64	FloTrac	SVV<10	
Van der Linden ⁴⁶	Vascular, 37	FloTrac	CI > 2.5	
Zeng ⁴⁷	Abdominal, 60	FloTrac	8 <svv<13< td=""></svv<13<>	
Zhang ⁴⁸	Thoracic, 60	FloTrac	SVV<11, CI > 2.5	
Zheng ⁴⁹	Abdominal, 60	FloTrac	SVV<12, SVI>35, CI>2.5	

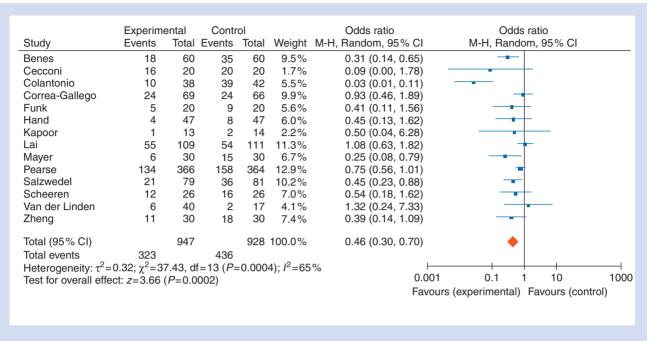
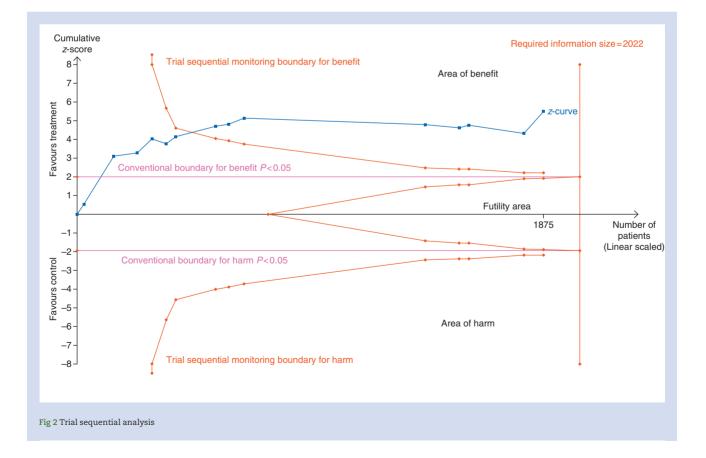
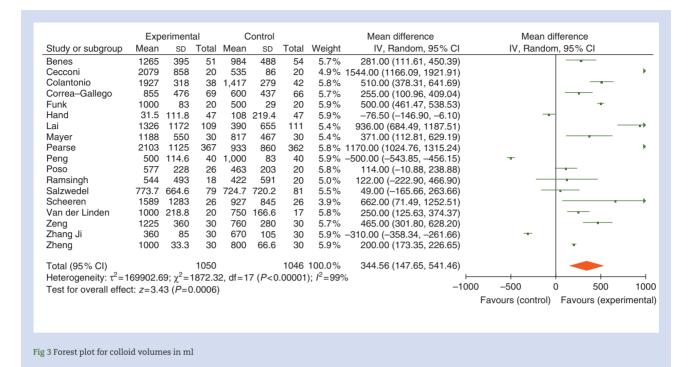


Fig 1 Forest plot for postoperative morbidity (defined as the proportion of patients who developed one or more complications)





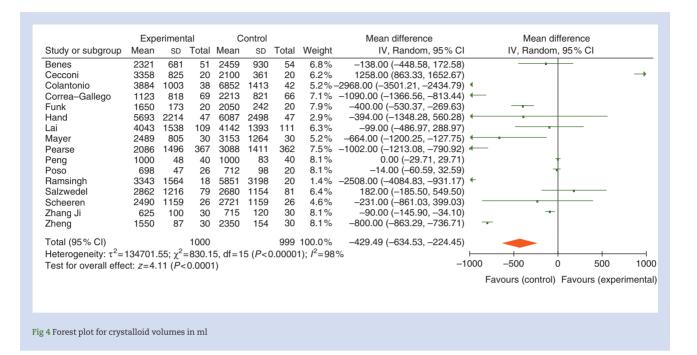
when using the fixed-effects model (see Supplementary data Table S4). Variability of fluid volume was 34% in the GDT group and 33% in the control group (P=0.98) (Fig. 6).

Discussion

Our meta-analysis shows that the use of GDT with uPC methods is associated with a significant decrease in postoperative morbidity. This effect was mainly related to a decrease in infectious, cardiac and abdominal complications. It was associated with a non-significant reduction in hospital length of stay (-1.35 days, P=0.06) and no change in mortality (2.6% vs 2.8%). These findings are consistent with results from previous meta-analyses where more complex and/or more invasive cardiac output monitoring technologies, such as pulmonary or transpulmonary thermodilution, oesophageal Doppler and saline- or lithiumcalibrated pulse contour methods, ^{1–3 52 53} were used for GDT. On one hand, the accuracy and precision of uPC methods has been questioned in patients with high cardiac output and low systemic vascular resistance, such as patients with septic shock receiving vasopressors or patients undergoing liver transplantation.^{19 54} On the other hand, uPC methods have been shown to be more reliable in patients undergoing surgery,⁵⁵ in particular to track haemodynamic changes during fluid loading.⁵⁶ A recent sub-analysis of the large EUSOS study¹⁶ showed that uPC methods became in a few years the preferred choice of European clinicians for haemodynamic monitoring in non-cardiac surgical patients. Our meta-analysis is the first to suggest that, despite their limitations, uPC methods are useful to guide haemodynamic management and improve the postoperative outcome of patients undergoing major surgery.

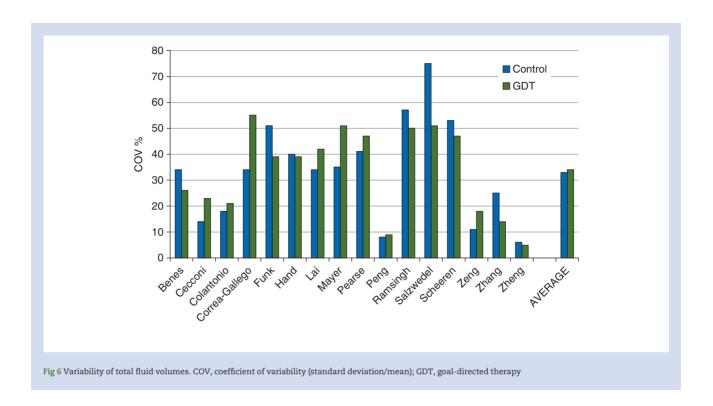
Another important and new finding of our meta-analysis is that the total volume of fluid administered during the study period did not increase with the use of GDT. Patients received more colloids, but less crystalloids, so that the total volume of fluid was not significantly different between the control and the GDT group when using the random-effects model, and lower with GDT when using the fixed-effects model. This finding goes against the perception or the fear that using haemodynamic optimization protocols may be associated with excessive fluid administration. As a matter of fact, GDT protocols significantly evolved over time from the maximization of oxygen delivery proposed almost 30 years ago by Shoemaker and colleagues⁸ to more individualized approaches restricting fluid administration to fluid responders only.44 45 Because we focused on studies done with uPC methods, which were not available on the market before 2006, our analysis was limited to recent GDT studies (all were published from 2008). This also contributes to the originality of our work, which gives information about the effects of GDT in the era of individualized fluid titration rather than at the time of aggressive haemodynamic strategies.

Recent studies have reported a large intra- and inter-variability of perioperative fluid volumes. In a database study involving 5912 patients undergoing abdominal surgery, Lilot and colleagues²³ showed that a patient weighing 75 kg who had a 4h procedure with 400 ml blood loss and 1 ml kg⁻¹ h⁻¹ urine output may have received anything between 700 and 5400 ml of crystalloid depending on the anaesthesiologist in charge. Thacker and colleagues²² also reported a large variability in fluid volumes in over 650000 patients undergoing abdominal and orthopaedic surgeries, and suggested that it may be responsible, at least in part, for postoperative adverse events. It has therefore been suggested that reducing variability with GDT protocols may be a way to improve postoperative outcome.^{22 26} Our meta-analysis suggests that the use of GDT is not associated with a reduction in the variability of fluid volume. Actually, the easiest way to decrease variability would be to give the exact same volume of fluid with an infusion pump to all surgical patients (e.g. 5 or $10 \text{ ml kg}^{-1} \text{ h}^{-1}$). However, it is unlikely that one size could fit all patients because they have different fluid needs, depending on what happened during the preoperative



	Expe	rimental		Co	ontrol			Mean difference	Mean difference
tudy or subgroup	Mean	SD	Total	Mean	SD	Total	Weight	IV, Random, 95% CI	IV, Random, 95% Cl
enes	3.586	0.926	51	3.444	1.183	54	7.1%	0.14 (-0.26, 0.55)	
ecconi	6.032	1.388	20	2.635	0.36	20	6.3%	3.40 (2.77, 4.03)	
olantonio	5.812	1.244	38	8.269	1.452	42	6.5%	-2.46 (-3.05, -1.87)	
orrea–Gallego	2	1.1	69	2.9	1	66	7.2%	-0.90 (-1.25, -0.55)	
unk	2.785	1.085	20	3.008	1.532	20	5.6%	-0.22 (-1.05, 0.60)	
and	5.887	2.291	47	6.319	2.519	47	5.1%	-0.43 (-1.41, 0.54)	
ai	5.369	2.27	109	4.532	1.525	111	6.7%	0.84 (0.33, 1.35)	
layer	4.528	2.317	30	4.494	1.561	30	5.0%	0.03 (-0.97, 1.03)	
earse	4.189	1.953	367	4.022	1.643	362	7.4%	0.17 (-0.09, 0.43)	-
eng	1.85	0.169	40	2.225	0.175	40	7.7%	-0.38 (-0.45, -0.30)	*
amsingh	4.082	2.044	18	6.845	3.893	20	2.5%	-2.76 (-4.71, -0.81)	
alzwedel	3.854	1.954	79	3.771	2.827	81	5.9%	0.08 (-0.67, 0.83)	
cheeren	4.477	2.107	26	4.528	2.387	26	4.2%	–0.05 (–1.27, 1.17)	
eng	2.732	0.488	30	3.135	0.346	30	7.5%	-0.40 (-0.62, -0.19)	
hang Ji	0.985	0.135	30	1.385	0.35	30	7.6%	-0.40 (-0.53, -0.27)	T
heng	2.65	0.133	30	3.95	0.221	30	7.7%	-1.30 (-1.39, -1.21)	*
Total (95 % Cl) 1004 1009 100.0 % -0.22 (-0.59, 0.15)								•	
eterogeneity: $\tau^2 = 0$.	47; $\chi^2 = 55$	0.26, df	=15 (<i>P</i> <	0.00001); I ² =97	7%		-	-+ + + + +
est for overall effect	t: z=1.15 (P=0.25)						-4 -2 0 2 4
									Favours (control) Favours (experime

period (fasting, bowel preparation, carbohydrate drinks, preinduction fluid load), on the clinical context (dehydration, haemorrhage, congestive heart failure, chronic renal failure) and on surgical blood loss. Therefore, variability in fluid volume may simply reflect the ability of clinicians to adapt fluid administration to individual patient needs. In 6248 patients undergoing high risk surgery, Kim and colleagues⁵⁷ recently showed that over 90% of the variability in crystalloid administration can be attributed to patient factors whereas only 10% was due to factors at the level of the care provider. One may then hypothesize that GDT helps clinicians to give the right amount of fluid to the right patients at the right time, without necessarily modifying the average amount of fluid given to a patient population (Fig. 5) nor fluid volume variability (Fig. 6). The importance of timing and individual titration intuitively makes sense, but our study design does not allow us to confirm nor reject this hypothesis. Finally, because crystalloids and colloids have different physiological and side effects, GDTinduced changes in the crystalloid/colloid ratio may have also played a role in the observed reduction in postoperative morbidity.



The main limitation of our meta-analysis is the high heterogeneity of studies analysed, which does not allow us to draw definitive conclusions. This limitation is common to most meta-analysis previously published on GDT, and is explained by differences between patient populations, surgical procedures, GDT protocols and definitions of postoperative complications from one RCT to the other. Interestingly, our results were comparable with the random-effects and the fixed-effects models for all but one outcome variables, which reinforced their validity. In addition, the trial sequential analysis, which has been proposed to unmask false-positive results in meta-analysis, 50 51 suggested firm evidence regarding the effects of GDT on postoperative morbidity. The average variability in fluid volume reported in our meta-analysis (around 33%) was lower than variabilities (55% and 70%) reported in recent studies.23 57 In one of the RCTs we analysed,44 the variability was very high in the control group (>70%) and decreased (below 50%) with GDT (Fig. 6). Therefore, we cannot exclude an impact of GDT on variability of fluid volume when variability is high (or at least higher than in the present meta-analysis). Another limitation is the fact that our meta-analysis did not take into account the increasing use of enhanced recovery programmes⁵⁸ and/or of the zero fluid balance approach.⁵⁹ By rationalizing and improving quality of perioperative surgical care, both strategies may diminish the marginal value of GDT.^{59 60} As a matter of fact they were not used in most of the trials we analysed. Therefore, a large study would be welcome to assess the effects of GDT with uPC methods in the era of enhanced recovery programmes and zero fluid balance.

Conclusions

Our meta-analysis shows that GDT with uPC methods is associated with a significant decrease in postoperative morbidity. It also shows that GDT with uPC does not increase the volume of fluid administered, and does not decrease the variability of fluid volumes. Therefore, our findings support the notion that uPC methods are useful to guide haemodynamic therapy during the perioperative period, and that studies are needed to clarify by which mechanisms GDT improves postoperative outcome.

Authors' contributions

Contributed equally to the design of the study, interpretation of results and preparation of the manuscript: all authors. Performed the statistical analysis and created the forest plots: M.T.G.

Supplementary material

Supplementary material is available at British Journal of Anaesthesia online.

Acknowledgements

We would like to thank the following authors for providing additional information and data from their original studies: Jan Benes, Mary Fischer, Duane Funk, William Hand, Gary Minto, Tomi Myrburg, Rupert Pearse, Davinder Ramsingh and Thomas Scheeren.

Declaration of interest

F.M. has been a consultant for Pulsion Medical Systems and an employee of Edwards Lifesciences. He is the founder and director of MiCo Sarl, a Swiss consulting firm. M.T.G. and N.B. have nothing to declare.

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Handling editor: J. G. Hardman